

The application of sampling theory in bauxite protocols

Teoria da amostragem aplicada aos protocolos de bauxita

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Abstract

The applications of Gy's formula are appropriate for calculating variances of the fundamental sampling error (FSE) at any stage of the sampling protocol and before samples are collected as well. However, the formula can be inaccurate because general factors are used to estimate the ore characteristic. In order to allow the calculation of fundamental sampling error and the minimum representative sample masses without using Gy's factors, there are experiments to calibrate the sampling parameters, namely: the heterogeneity test (HT); the sampling tree experiment (STE) and the segregation free analysis (SFA). The present work describes the experimental procedure for the three calibration methods using bauxite and shows the correlation between them.

Keywords: *sampling protocols; fundamental sampling error; bauxite.*

Resumo

A fórmula de Gy é apropriada para calcular a variância do erro fundamental de amostragem (FSE), em qualquer estágio do protocolo de amostragem e, mesmo, antes de a amostra ser coletada. Entretanto a fórmula pode ser inaccurada por utilizar fatores aproximados para todos os tipos de minérios. Para calcular a variância do erro fundamental de amostragem e as massas mínimas das amostras sem usar os fatores de Gy, existem experimentos para calibrar os parâmetros de amostragem, a saber: teste de heterogeneidade (HT); experimento da árvore (STE) e análise de segregação livre (SFA). O presente trabalho descreve os procedimentos experimentais para a bauxita, utilizando esses três métodos de calibração, e apresenta a correlação entre eles.

Palavras chave: *Protocolos de amostragem, erro fundamental de amostragem, bauxita.*

1. Introduction

Sampling theory of broken material ores cannot be mentioned without a direct reference to Gy's fundamental contributions (Gy, 1998). According to Minnnett

(2007) apud François-Bongarçon (1998), the fundamental sampling error (FSE) is the smallest achievable residual average error, because of the physical and chemi-

cal composition as well as the particle size distribution. FSE can be determined by Gy's formula, applicable to all kinds of ore (Equation 1):

$$s_{FSE}^2 = c f g l dN^3 \left(\frac{1}{M_s} - \frac{1}{M_L} \right) \quad (1)$$

Where M_s is the mass of selected sample, M_L is the mass of the original lot, c is the mineralogical factor, f is the shape

factor, g is the granulometric factor, l is the liberation factor and dN is the "top size". Furthermore, the formula allows

the calculation of the minimum sample mass to represent a lot for a maximum variance. The experiments developed to

estimate the sampling constants K and alpha of Equation 2 are described as

follows, where K is a specific constant to a given ore type at a given grade, and

alpha, is the exponent of original Gy's cubic formula

$$s^2_{FSE} = K d_N^\alpha \left(\frac{1}{M_s} - \frac{1}{M_L} \right) \quad (2)$$

2. Methodology

Heterogeneity test (HT)

Initially, a 500 kg sample representing the typical aluminium ore was selected at the mine, then crushed and screened in four size fractions: “-38.1 mm +25.4 mm”; “-25.4 mm +12.7 mm”;

“-12.7 mm +6.3 mm”; “-6.3 mm +1.2 mm”. For each size fraction, 50 groups of at least 50 fragments each, collected one by one at random, were selected. Once samples were weighted and analyzed, the

mass M_q and grade a_q for each group of fragments, as well as the average masses M_Q and weighted average grades a_Q were calculated according to Equation 3 (Koyama et al., 2010).

$$a_Q = \frac{1}{M_Q} \sum a_q M_q \quad (3)$$

The estimated constant factor of constitutional heterogeneity, EST IHL, was calculated according to Equation

4, for each of the four size fractions. These factors generate a regression line calculated using the four resulting

points, which allows estimating the parameters K and alpha of Equation 2 (Pitard, 1993).

$$EST\ IHL = g \sum \frac{(a_q - a_Q)^2}{a_Q^2} \cdot \frac{M_q^2}{M_Q} \quad (4)$$

The heterogeneity test allows the experimental calibration of the sam-

pling constants K and alpha, which must be calibrated for each particular

ore at a particular grade.

Sampling tree experiment (STE)

The sampling tree experiment proposed by François-Bongarçon (Minnitt (2007) apud François-Bongarçon 1995 and 1998) follows a detailed analysis of the application of Gy's formula to determine the fundamental sampling error. A representative 60 kg sample was collected at the mine. The primary crushing stage using a jaw crusher reduces the ore nominal size to a uniform 95% passing 25.4 mm. One quarter of the total lot is split out

and forms the first nominal size fraction. The remainder of the ore is then crushed to 95% passing in 6.3 mm and the lot is then split into three equal portions. One of these portions is set aside and constitutes the second nominal size fraction. The remaining two fractions are recombined, crushed to 95% passing 2.0 mm and split into two portions of equal mass. One portion is set aside as the third nominal size fraction and the remaining one, crushed to

95% passing 0.1 mm, composes the fourth nominal size fraction. This procedure generates four portions of more or less equal mass, about 15 kg. Using a riffle splitter, each of these portions is now split into 32 sub-samples per fraction. All samples are weighted and chemically analysed. According to Minnitt et al. (2007), the form of Gy's equation can be changed by taking the logarithms of both sides, transposing to Equation 5.

$$\ln(s^2_{FSE} M_s) = \alpha \ln(d_N) + \ln(K) \quad (5)$$

Segregation free analysis (SFA)

Both the heterogeneity test and the sample tree experiment suffer from weaknesses that cast suspicion on the derived values for K and alpha. Minnitt

et al. (2011) proposed a new test called segregation free analysis (SFA). The nominal (average) fragment size, d_N , for a fragment passing between two screens

is given by Equation 6, where USS is the upper screen size and LSS is the lower screen size.

$$d_N = \sqrt[3]{\frac{USS^3 + LSS^3}{2}} \quad (6)$$

The material, 300 kg of crushed bauxite collected from many cuts in the sampling station, was classified at 14 size fractions: “-38.1 mm +25.4mm”, “-25.4 mm +12.7mm”; “-12.7 mm +6.3 mm”; “-6.3 mm +1.4 mm”; “-1.4 mm

+0.84 mm”; “-0.84 mm +0.60 mm”; “-0.60 mm +0.5 mm”; “-0.5 mm +0.3 mm”; “-0.3 mm +0.23 mm”; “-0.23 mm +0.15 mm”; “-0.15 mm +0.10 mm”; “-0.10 mm +0.07 mm”; “-0.07 mm + 0.05 mm”; “-0.05 mm + 0.03 mm”.

Each size fraction was split using a riffle splitter into a series of 32 samples and sent to the chemical laboratory. The variance of the sample series associated with each fragment size was calculated and plotted to produce a straight line

which can be used to calculate K , α and liberation size, d_L . Provided that the screen sizes are chosen such that the

ratio $r = d_{MAX}/d_{MIN}$ is reasonably consistent, all points representing different d_{MAX} values on the calibration curve will

$$s^2_{FSE} = c f g' l d^3 \left(\frac{1}{M_S} - \frac{1}{M_L} \right) \quad (7)$$

The granulometric factor g is replaced here by g' , where $g' = g'(r)$ is read off the curve of granulometric factor

g' vs ratio $r = d_{MAX}/d_{MIN}$, for closely screened materials.

The liberation factor is the transi-

represent the formula for the relative variance given by Equation 7.

tion between the liberated, calculable variance, to the non-liberated one given by Equation 8.

$$l = \left(\frac{d_L}{d_{MAX}} \right)^{(3-\alpha)} \quad (8)$$

Equation 9 results from substituting Equation 8 in Equation 7 and

taking the log on both sides. The linear regression built using the 14 resulting

points gives the sampling parameters K and α .

$$\ln(S^2_{rel} M_S) = \alpha \ln(d_{MAX}) + \ln(K) \quad (9)$$

3. Results and discussion

Heterogeneity test (HT)

EST IHL was calculated using Equation 4. Table 1 shows the results

of EST IHL for each fragment size (shown in cm). The granulometric

factor for calibrated material is 0.55 (Gy, 1998).

d_N	3.32	2.09	1.05	0.5
EST IHL	0.823	0.337	0.149	0.008

Table 1

EST IHL results for corresponding nominal size

Figure 1 shows the correlation between the nominal size and $EST IHL$. The equation of the power trendline

($y=c \cdot x^b$) built for the four fragment sizes gives the sampling parameters K (c of the trendline equation) and α

(b of the trendline equation).

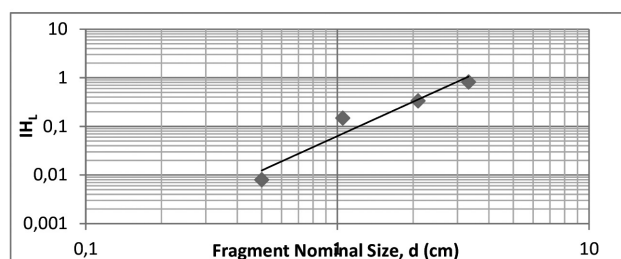


Figure 1

Correlation between IHL and the fragment nominal size

According to Figure 1, the estimates

for K and α are 0.0638 and 2.34

respectively.

Sampling tree experiment (STE)

The alumina content and the sample masses were submitted to a

statistical treatment as described by Minnitt et al. (2007) and are presented

in Table 2.

Serie	Total Mass	Average Grade	CV	Relative VAR	Std. Rel VAR	$\ln(d_N)$	$\ln(\sigma^2 \times M_s)$
1	477.3	48.17	0.0254	0.0006	0.0312	0.9322	2.7005
2	460.48	47.86	0.0126	0.0002	0.0076	-0.4541	1.2503
3	459.96	47.62	0.0195	0.0004	0.0181	-1.6094	2.1219
4	407.03	48.06	0.0067	0	0.0022	-4.5469	-0.1191

Table 2

STE statistical results

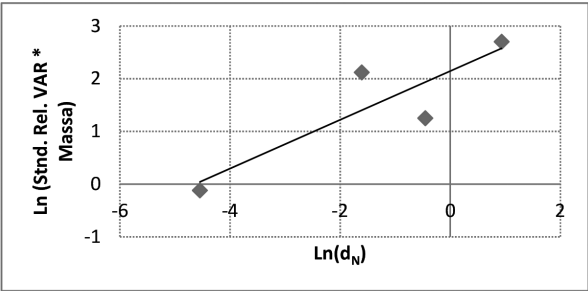
Where relative variance is CV^2 (coefficient of variation), standard

relative variance (σ^2) is the product between average grade and relative vari-

ance, and M_s is the sample average mass. Four fragment sizes were used to build

the regression line (Figure 2) according to Equation 5.

Figure 2
K and alpha calibration curve for a 47.8% Al₂O₃ ore.



The exponential function of the intercept on the y-axis provides an

estimate of K. According to the sampling tree experiment, the estimates

for K and alpha are 8.54 and 0.46 respectively.

Segregation free analysis (SFA)

The alumina content and the sample masses for the 14 size fractions were submitted to a statistical treatment as

described by Minnitt et al. (2011) and are presented in Table 3. The relative variance is CV² (coefficient of variation), standard

relative variance (σ^2) is the product between average grade and relative variance and Ms is the sample average mass.

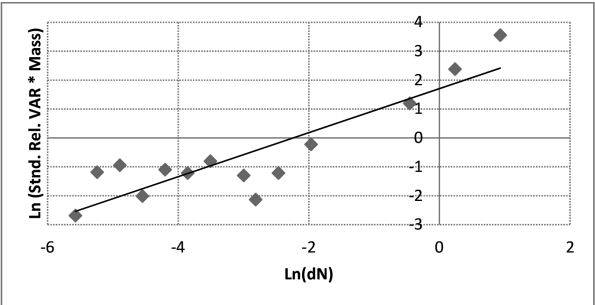
Table 3
SFA statistical results

Serie	Total Mass	Average Grade	CV	Relative VAR	Std. Rel VAR	Ln(dN)	Ln($\sigma^2 \times M_s$)
1	360.07	50.98	0.04	0.0019	0.097	0.932	3.556
2	168.19	51.17	0.04	0.0013	0.064	0.239	2.381
3	176.61	48.77	0.02	0.0004	0.019	-0.454	1.197
4	11w3.37	50.51	0.01	0.0001	0.007	-1.966	-0.223
5	59.80	47.74	0.01	0.0001	0.005	-2.465	-1.215
6	27.58	49.51	0.01	0.0001	0.004	-2.813	-2.132
7	43.29	49.66	0.01	0.0001	0.006	-2.996	-1.300
8	26.34	50.32	0.02	0.0003	0.017	-3.507	-0.803
9	38.11	49.78	0.01	0.0002	0.008	-3.854	-1.213
10	40.08	50.48	0.01	0.0002	0.008	-4.200	-1.097
11	46.21	53.14	0.01	0.0001	0.003	-4.547	-2.010
12	34.32	50.83	0.01	0.0002	0.011	-4.893	-0.942
13	40.82	50.53	0.01	0.0001	0.007	-5.240	-1.184
14	39.64	50.03	0.01	0.0000	0.002	-5.573	-2.686

The 14 fragment sizes were used

to build the regression line on Figure 3, according to Equation 9.

Figure 3
K and alpha calibration curve for a 50.2% Al₂O₃ ore



According to SFA, the estimates for K and alpha are 5.49 and 0.76 respectively.

Comparison between sampling protocols using HT, STE and SFA

Table 4 shows the values of Gy's factors (Pitard, 1993) for Juruti's bauxite according to Equation 1.

Factor	c	f	g	di	Al ₂ O ₃	λM	λG	IH _L
Values	3.6	0.5	0.25	0.00036	48.1%	3.98	2.65	0.0085 d ^{2.5}

Table 5 shows the sampling and sample preparation protocol for Juruti's mine and the relative standard deviation

of the fundamental sampling error (last 4 columns, in %), s_{FSE} , of each stage using Gy's factors (Gy,1998) as well as the

Step	M _L (g)	MS (g)	d _N (cm)	s _{FSE} GY	s _{FSE} HT	s _{FSE} STE	s _{FSE} SFA
Primary Sampling	1.5E+10	15000	5	0.5628	1.3461	3.4618	3.5328
Crushing	15000	15000	2.54	0.0000	0.0000	0.0000	0.0000
Primary Quartering	15000	2500	2.54	0.5397	1.3632	6.6183	6.1023
Grinding	2500	2500	0.2	0.0000	0.0000	0.0000	0.0000
Secondary Quarterly	2500	200	0.2	0.0836	0.2592	13.6578	8.6044
Pulverization	200	200	0.015	0	0	0	0
TOTAL				0.784 %	1.933 %	15.567 %	11.125 %

Considering Equation 2, a maximum relative standard deviation of the fundamental error of 3% and a high value

for the initial mass M_L, the minimum sample mass to represent the primary sample at the mine, where d_{MAX} is 5 cm,

Parameter	GY	HT	STE	SFA
d _{MAX} (cm)	5	5	5	5
M _s (g)	530	3.063	19.895	20.728
d _{MAX} (cm)	1	1	1	1
M _s (g)	9	71	9.486	6.099

4. Conclusions

This paper presented a comparison of different methods to estimate the variance of the fundamental sampling error, minimum sample masses and the sampling parameters by different methods. An appropriate sampling protocol for any type of ore requires experimental test work in order to determine the deposit's heterogeneity. The heterogeneity test (HT) has been used for several decades and has proved to be an excellent method to calculate the minimum representative sample masses and to optimize sampling protocols during the beginning of mining projects and during operation as well. However, the new proposed tests show significant differences

which should be taken into account. The comparison between the results of STE and SFA shows the relationship between the fundamental sampling error and the grouping and segregation error. SFA was developed aiming to eliminate the residual grouping and segregation variance which biases the calibration curve and should present smaller variances than STE. However, for higher diameters, SFA presents higher values of s_{2SFE}, suggesting that the fundamental sampling error makes the most important role for coarser particles. Based on the results, the company decided to adopt the segregation free analysis for defining its sampling protocols due to two major

Table 4

Factors used in Gy's relative variance equation

results of HT, STE, SFA.

Table 5

Sampling protocol comparing the three heterogeneity experiments

can be calculated. The last line of table 6 shows the same calculations for a d_{MAX} of 1 cm.

Table 6

Minimum sample masses (Ms) for the primary sample at the mine

reasons: (1) the SFA proved to reduce the grouping and segregation error and (2) the SFA presented higher minimum sample masses than the heterogeneity test, making this method safer than the HT with respect to the representativeness of samples. The results showed that optimizing sampling protocols is not a simple task and more than one experimental test should be performed to confirm what the minimum sample masses should be. This paper is a study for bauxite and may be not applicable to other types of ore. The authors highly recommend other companies to compare the different approaches and define the most suitable for their operations.

5. Acknowledgments

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